Decision Theory and Performance Analysis

EE564: Digital Communication and Coding Systems

Keith M. Chugg Spring 2017 (updated 2020)

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Course Topic (from Syllabus)

- Overview of Comm/Coding
- Signal representation and Random Processes
- Optimal demodulation and decoding
- Uncoded modulations, demod, performance
- Classical FEC
- Modern FEC
- Non-AWGN channels (intersymbol interference)
- Practical consideration (PAPR, synchronization, spectral masks, etc.)

Detection/Demod Topics

- Maximum A Posteriori decision rule for vector-AWGN channel
- Exact performance for binary modulations
- Minimum distance decision rule for M-ary modulation over AWGN
- Performance bounds
- Continuous time model
	- Likelihood functional, sufficient statistics
- Average and generalized likelihood
	- Phase non-coherent demodulation
	- Soft-out demodulation

Figure 1.1. The set-up for general decision problems considered.

Decision Problem 1.1.1 Decision Pro 1.1.1 The Bayes Decision Rule The finite set of coefficients *{C*(*Am, Hi*)*}m,i* specifies the cost of taking A Bayes decision rule minimizes the average Bayes risk *R*(*d*) over all action *A^m* with *Hⁱ* occurring. These coefficients relate the Bayes risk to $\overline{}$ *i* P

development.
Development Bayes risk for

Bayes risk for
decision rule d
$$
R(d) = \int_{\mathcal{Z}} p_{\mathbf{z}(\zeta)}(\mathbf{z}) \left[\sum_{m} d(A_m | \mathbf{z}) C(A_m | \mathbf{z}) \right] d\mathbf{z}
$$
(1.1)

ciated with taking action *Am*, given z(ζ) = z, averaged over the source $\frac{18}{15}$ ciated with taking action *Am*, given z(ζ) = z, averaged over the source Cost for taking these best actions in any way. One can always the can al rule is generally not unique because of "tie" conditions where two or "tie" conditions where two or "tie" cond
Ties where two or "tie" conditions where two or "tie" conditions where two or "tie" conditions where two or "t

Cost for taking
\naction m given
\nobservation
\n
$$
C(A_m | \mathbf{z}) = \sum_i C(A_m, H_i) p_{H(\zeta)|\mathbf{z}(\zeta)}(H_i | \mathbf{z})
$$
\n(1.2)

Bayes decision rule

sion rule Bayes action = arg min
$$
C(A_m | \mathbf{z})
$$
 (1.3)

$$
\text{APP factoring} \qquad p_{H(\zeta)|\mathbf{z}(\zeta)}(H_m|\mathbf{z}) = \frac{p_{\mathbf{z}(\zeta)|H(\zeta)}(\mathbf{z}|H_m)p_{H(\zeta)}(H_m)}{p_{\mathbf{z}(\zeta)}(\mathbf{z})} \qquad (1.4a)
$$
\n
$$
\equiv p_{\mathbf{z}(\zeta)|H(\zeta)}(\mathbf{z}|H_m)p_{H(\zeta)}(H_m) \qquad (1.4b)
$$

MAP Decision Rule

MAP is special case of Bayesian Decision Rule

The *Maximum A-Posteriori Probability (MAP)* decision rule is the special case of the Bayes rule when A_m corresponds to deciding that H_m is true and $C(A_m, H_i) = 1 - \delta_{m-i}$. This may be seen by substituting these cost coefficients into (1.2) and noting that

$$
C(A_m|\mathbf{z}) = \sum_{i \neq m} p_{H(\zeta)|\mathbf{z}(\zeta)}(H_i|\mathbf{z}) = 1 - p_{H(\zeta)|\mathbf{z}(\zeta)}(H_m|\mathbf{z}) \tag{1.5}
$$

MAP rule from P_errror Expression 2 *e|z[|]*

 $\pi_0 = \pi_1 = 0.5$ $f_{z(u)}(z|\mathcal{H}_0) = \frac{1}{2}$ 2 $e^{-|z|}$ $f_{z(u)}(z|\mathcal{H}_1) = \frac{1}{2}$ 2 $\mathcal{N}(z; -2; 1) + \frac{1}{2}$ 2 $\mathcal{N}(z; +2; 1)$ $\frac{1}{|z|}$ $\frac{1}{|z|}$

 $\epsilon \in \mathcal{R}: I_0(z) >$ *M*!1 $I_0(z) > I_1(z) \}$ *M*!1 lim Design Z0 and Z1: $\mathcal{Z}_0 = \{ z \in \mathcal{R} : I_0(z) > I_1(z) \}$

MAP Rule for Vector-AWGN Channel

 $\overline{}$

$$
\mathcal{H}_m: \quad \mathbf{z}(u) = \mathbf{s}_m + \mathbf{w}(u) \qquad (D \times 1)
$$

$$
P(\mathcal{H}_m|\mathbf{z}) = \frac{f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m)\pi_m}{f_{\mathbf{z}(u)}(\mathbf{z})}
$$

\n
$$
\equiv f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m)\pi_m
$$

\n
$$
= \mathcal{N}_D(\mathbf{z}; \mathbf{s}_m; (N_0/2)\mathbf{I})\pi_m
$$

\n
$$
= \frac{\pi_m}{(\pi N_0)^{D/2}} \exp\left[\frac{-1}{N_0} ||\mathbf{z} - \mathbf{s}_m||^2\right]
$$

\n
$$
\equiv \pi_m \exp\left[\frac{-1}{N_0} ||\mathbf{z} - \mathbf{s}_m||^2\right]
$$

$$
\begin{pmatrix}\n\max_{m} P(\mathcal{H}_{m}|\mathbf{z}) & \Longleftrightarrow & \min_{m} - \ln (P(\mathcal{H}_{m}|\mathbf{z}) \\
& \Longleftrightarrow & \min_{m} \left[-\ln(\pi_{m}) + \frac{1}{N_{0}} ||\mathbf{z} - \mathbf{s}_{m}||^{2} \right] \\
& \Longleftrightarrow & \min_{m} ||\mathbf{z} - \mathbf{s}_{m}||^{2} \quad \left(\text{when } \pi_{m} = \frac{1}{M} \right)\n\end{pmatrix}
$$

Other Rules (MAP Special Cases)

Maximum Likelihood (ML):

 $\max_m f(\mathbf{z}|\mathcal{H}_m)$

Minimum Distance:

Min. Euclidean (squared) distance:

m $d(\mathbf{z}, \mathbf{s}_m)$

 $\frac{\min}{m} \|\mathbf{z} - \mathbf{s}_m\|^2$

$M=2$

Maximum Likelihood (ML): $\frac{\mu_1}{\lambda}$ *>* **
≤** H_0 $f(\mathbf{z}|\mathcal{H}_0)$ Minimum Distance: $\frac{\mu_1}{\lambda}$ *> <* H_0 $d(\mathbf{z}, \mathbf{s}_1)$ Min. Euclidean (squared) distance: $\frac{\mu_1}{\lambda}$ *>* **
∕** μ_0 $\|\mathbf{z}-\mathbf{s}_1\|^2$

MAP reduces to ML when a priori probabilities are uniform

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Other Rules (MAP Special Cases)

Binary MAP Decisions (equal priors)

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$$
\begin{aligned} \textbf{Performance of Binary MAP Decisions (equal priors)}\\ P(\mathcal{E}|\mathcal{H}_0) &= \Pr\left\{(s_1-s_0)^t\mathbf{z}(u) > \frac{\|\mathbf{s}_1\|^2 - \|\mathbf{s}_0\|^2 - N_0\ln(\pi_1/\pi_0)}{2}|\mathcal{H}_0\right\} \\ &= \Pr\left\{(s_1-s_0)^t(s_0+\mathbf{w}(u)) > \frac{\|\mathbf{s}_1\|^2 - \|\mathbf{s}_0\|^2 - N_0\ln(\pi_1/\pi_0)}{2}\right\} \\ &= \Pr\left\{(s_1^t s_0 - \|\mathbf{s}_0\|^2) + (s_1-s_0)^t\mathbf{w}(u) > \frac{\|\mathbf{s}_1\|^2 - \|\mathbf{s}_0\|^2 - N_0\ln(\pi_1/\pi_0)}{2}\right\} \\ &= \Pr\left\{(s_1-s_0)^t\mathbf{w}(u) > \frac{\|\mathbf{s}_1\|^2 + \|\mathbf{s}_0\|^2 - 2s_1^t s_0 - N_0\ln(\pi_1/\pi_0)}{2}\right\} \\ &= \Pr\left\{V(u) > \frac{1}{2}\left[\|\mathbf{s}_1 - \mathbf{s}_0\|^2 - N_0\ln(\pi_1/\pi_0)\right]\right\} \end{aligned}
$$

 $\mathbb{E}\left\{V(u)\right\}=0$

$$
\sigma_V^2 = \frac{N_0}{2} ||\mathbf{s}_1 - \mathbf{s}_0||^2
$$

Performance of Binary MAP Decisions (equal priors)

$$
P(\mathcal{E}|\mathcal{H}_0) = Q\left(\sqrt{\frac{d^2}{2N_0}}\right)
$$

$$
d^2 = ||\mathbf{s}_1 - \mathbf{s}_0||^2 \quad (\pi_1 = \pi_0)
$$

$$
\begin{pmatrix}\n\frac{d}{2} \\
\frac{d}{2} \\
\frac{d}{2} \\
\frac{d}{2}\n\end{pmatrix}
$$
\n\nContours of $f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_0)$

 $P(\mathcal{E}) = P(\mathcal{E}|\mathcal{H}_0)\pi_0 + P(\mathcal{E}|\mathcal{H}_1)\pi_1$

$$
= P(\mathcal{E}|\mathcal{H}_0)(1/2) + P(\mathcal{E}|\mathcal{H}_1)(1/2)
$$

$$
P(\mathcal{E}) = Q\left(\sqrt{\frac{d^2}{2N_0}}\right) = Q\left(\sqrt{\frac{\|\mathbf{s}_1 - \mathbf{s}_0\|^2}{2N_0}}\right)
$$

Note: not a function of dimension

Performance of Binary MAP Decisions (equal priors)

$$
\|\mathbf{s}_1 - \mathbf{s}_0\|^2 = E_1 + E_0 - 2\mathbf{s}_1^t \mathbf{s}_0
$$

$$
\rho = \frac{\mathbf{s}_1^{\mathbf{t}} \mathbf{s}_0}{\sqrt{E_1 E_0}}
$$

$$
\|\mathbf{s}_1 - \mathbf{s}_0\|^2 = 2E(1 - \rho) \quad \text{(equal energy)}
$$

$$
\rho = \frac{\mathbf{s}_1^{\mathbf{t}} \mathbf{s}_0}{E} \qquad \text{(equal energy)}
$$

best equal energy binary signals are **antipodal signaling**

$$
P(\mathcal{E}) = Q\left(\sqrt{\frac{2E}{N_0}}\right) \quad \text{(antipodal)}
$$

$$
P(\mathcal{E}) = Q\left(\sqrt{\frac{E}{N_0}}\right) \quad \text{(orthogonal, coherent)}
$$

Binary orthogonal signaling is 3 dB worse than antipodal

Performance of Binary MAP Decisions (equal priors)

Blue and red curves show results from previous slide

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MAP Decision Regions 0*,* 1*,...M* 1 based on some measurement z 2 *Z*. Such a rule introduces a partition of the Consider an arbitrary deterministic decision rule for deciding among *M* hypotheses *H^m* for *m* = *^P*(*E|Hm*) = Pr *{*z(*u*) ⁶² *^Zm|Hm}* ⁼ Pr *{*z(*u*) ² *^Z^c* Many useful bounds can be constructed by expressing *^Z^c ^m* in specific ways. For example, it is clear *^P*(*E|Hm*) = Pr *{*z(*u*) ⁶² *^Zm|Hm}* ⁼ Pr *{*z(*u*) ² *^Z^c*

decide *H^m* () z 2 *Z^m* (1) $\frac{1}{2}$ $\frac{1}{2}$ $\frac{1}{2}$ $\frac{2}{m}$ decide $\mathcal{H}_m \iff \mathbf{z} \in \mathcal{Z}_m$ $\det \mathcal{H}_m$ \Leftarrow \mathcal{Z}_m Rec ide \mathcal{H}_m = \Leftarrow \Rightarrow $\epsilon \geq m$

(global) (giopal)
decision $\mathcal{Z} = \mathcal{I}_{\mathbf{z}}$ region regions *^ZPW* regions *^ZPW* where *Z^m* $\begin{bmatrix} 8 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix}$ where *Z^m*

decision
region
$$
\mathcal{Z}_m = \{ \mathbf{z} : f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m) P(\mathcal{H}_m) > f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_j) P(\mathcal{H}_j) \ \forall \ j \neq m \}
$$

pairwise decision region (m over j)

decision region

\n
$$
\mathcal{Z}_m^{PW}(j) = \{ \mathbf{z} : f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m) P(\mathcal{H}_m) > f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_j) P(\mathcal{H}_j) \}
$$
\n(m over j)

$$
\mathcal{Z}_m = \bigcap_{j=0,j\neq m}^{M-1} \mathcal{Z}_m^{PW}(j)
$$

Comparing the definitions of the global decision region and the pairwise regions (i.e., (2) and (5))

In order for 'Hm' to be the decision, it must better than every other hypothesis in a pairwise test *j*_p and *in* for 4 *decision* it must bet *<i>P* (*z*) $\frac{1}{2}$ (*j*) $\frac{1}{2}$) $\frac{1}{2}$ (*p*) $\frac{1$ every other hypothesis in a pairwise test every for 'Hm' to be the decision, it must
were other hypothesis in a pairwise Let by pothesis in a pai st D
se te *m* to be the decis *j*₆
*j*6=0,*jp6*
*j*6=0,*jp6*
*j*6=0,jp ion, านรt *z*
z<i>z (*x*) (*m*) (*m*

The compliment of this region is obtained by applying $\overline{\mathcal{L}}$. This region is obtained by applying $\overline{\mathcal{L}}$

8-PSK Example Min. Distance Rule

4-PAM Example Min. Distance Rule

16-QAM Example Min. Distance Rule

$$
\mathcal{Z}_m^c = \bigcup_{j=0,j\neq m}^{M-1} \left[\mathcal{Z}_m^{PW}(j) \right]^c = \bigcup_{j=0,j\neq m}^{M-1} \mathcal{Z}_j^{PW}(m)
$$

$$
P(\mathcal{E}|\mathcal{H}_m) = \Pr \{ \mathbf{z}(u) \in \mathcal{Z}_m^c | \mathcal{H}_m \}
$$

$$
= \Pr \left\{ \mathbf{z}(u) \in \bigcup_{j=0, j \neq m}^{M-1} \mathcal{Z}_j^{PW}(m) \right\}
$$

$$
\text{union} \qquad \max_{j} P_{PW}(j|\mathcal{H}_m) \le P(\mathcal{E}|\mathcal{H}_m) \le \sum_{j=0, j \ne m}^{M-1} P_{PW}(j|\mathcal{H}_m)
$$

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union

Performance Bounds for M-ary is more generally applicable. In this case, and the this case, applicable the the expression in ϵ P ance Bounds for M $\overline{}$

$$
\sum_{m=0}^{M-1} P(\mathcal{H}_m) \left[\max_j P_{PW}(j|\mathcal{H}_m) \right] \le P(\mathcal{E}) \le \sum_{m=0}^{M-1} P(\mathcal{H}_m) \sum_{j \in \mathcal{N}_m}^{M-1} P_{PW}(j|\mathcal{H}_m)
$$

Performance Bounds: M-ary AWGN *PPW* (*j|Hm*)=Q @ 2*N*⁰ ^A *^d*2(*j, m*) = ^ks*^j* ^s*m*k² (13) *Notes on Performance Bounds v1.4 -* c *K.M. Chugg – October 5, 2015* 3 \mathcal{A} common special case for the application of the bounds developed above is that of a-priori equal \mathcal{A} likely signaling over an AWGN where μ and μ and μ and μ are μ and μ this case, the MAP detection rule is the MAP detection rule is the Minimum Distance rule and the pairwise error Performance Bounds: M-ary AWGN

equal priors

$$
\text{equal priors} \quad \overbrace{\left(\frac{1}{M} \sum_{m=0}^{M-1} \mathbf{Q} \left(\sqrt{\frac{d_{\min}^2(m)}{2N_0}} \right) \le P(\mathcal{E}) \le \frac{1}{M} \sum_{m=0}^{M-1} \sum_{j \in \mathcal{N}_m} \mathbf{Q} \left(\sqrt{\frac{d^2(j,m)}{2N_0}} \right)}_{j \neq m} \right)}
$$
\n
$$
d_{\min}^2(m) = \min_{j \neq m} d^2(j,m) \qquad d^2(j,m) = \|\mathbf{s}_j - \mathbf{s}_m\|^2
$$

2*N*⁰

M X1 *M* X1

*d*2(*j, m*)

*d*2 min = min *d*2 min(*m*) (16) Ing to combine terms with same diepting *M* com *Aook-keeping to combine terms with same distances* \mathbf{b} bound in (14) then simplifies to bound in (14) Book-keeping to combine terms with same distances

1.1 Special Cases for AWGN Channels for
1971 - Awgnesia Cases for AWGN Channels for AWGN Channels for AWGN Channels for AWGN Channels for AWGN Channel

$$
\sum_i \frac{K_i}{M} \mathrm{Q}\left(\sqrt{\frac{d_i^2}{2N_0}}\right) \leq P(\mathcal{E}) \leq \sum_i \frac{N_i(\{\mathcal{N}_m\})}{M} \mathrm{Q}\left(\sqrt{\frac{d_i^2}{2N_0}}\right)
$$

which inputs that at the error probability must decay probability of α α β Less tight, but very simple version: The global minimum distance of Less tight, but very simple version: bestight but very simple version: $\frac{1}{2}$ and $\frac{1}{2}$ $\frac{1}{2}$ and $\frac{1}{2}$ and $\frac{1}{2}$

$$
\frac{1}{M}\mathrm{Q}\left(\sqrt{\frac{d_{\min}^2}{2N_0}}\right)\leq P(\mathcal{E})\leq (M-1)\mathrm{Q}\left(\sqrt{\frac{d_{\min}^2}{2N_0}}\right)
$$

Performance Bounds: M-ary AWGN

$$
\frac{K_1}{M} \mathcal{Q}\left(\sqrt{\frac{d_{\min}^2}{2N_0}}\right) \le P(\mathcal{E}) \approx \frac{N_1(\{\mathcal{N}_m\})}{M} \mathcal{Q}\left(\sqrt{\frac{d_{\min}^2}{2N_0}}\right)
$$

performance is dominated by the minimum distance

Performance Bounds: Bit Error Probability *th* bit is decided in error. Specifically, let *^Bⁱ* be the event that *^P*(*Bi*) = *^P*(*Bi|E*)*P*(*E*) = *P*(*Bi|E*)*P*(*E*) (22) Performance Bounds: Bit Error Probability evaluate *P*(*Bi|E*), but we can consider the average bit error probability (i.e., averaged over bit Performance Bounds: Bit Frror Probability

$$
P(\mathcal{B}_i) = \frac{P(\mathcal{B}_i | \mathcal{E}) P(\mathcal{E})}{P(\mathcal{E} | \mathcal{B}_i)} = P(\mathcal{B}_i | \mathcal{E}) P(\mathcal{E})
$$

$$
P_b = \frac{1}{q} \sum_{i=0}^{q-1} P(\mathcal{B}_i) = \frac{1}{q} \sum_{i=0}^{q-1} P(\mathcal{B}_i | \mathcal{E}) P(\mathcal{E})
$$

$$
\frac{1}{q}P(\mathcal{E}) \le P_b \le P(\mathcal{E})
$$
\n
$$
\frac{1}{q}B_L(\mathcal{E}) \le P_b \le B_U(\mathcal{E})
$$
\nlower bound on
symbol error
symbol error
symbol error

*k*1

 a symbol error probability

*k*1

pper bound on symbol error probability where *BL*(*E*) and *B^U* (*E*) are upper and lower bounds on *P*(*E*), respectively. Note that this approach where *B*_Z(*E*) and *B*_{*U*} (*E*) and *B*_{*Z*(*E*), *B*^{*Z*} and *B*^{*Z*}), *B*^{*Z*} and *E*), *B*}

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	- *• Performance of common M-ary modulations*
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Performance Bounds: MPSK

$$
d^{2}(m, n) = \left| \sqrt{E_{s}} e^{j\frac{2\pi}{M}m} - \sqrt{E_{s}} e^{j\frac{2\pi}{M}n} \right|^{2}
$$

\n
$$
= E_{s} \left(1 + 1 - 2\Re \left\{ e^{j\frac{2\pi}{M}(m-n)} \right\} \right)
$$

\n
$$
= 2E_{s} (1 - \cos \left(\frac{2\pi}{M}(m-n) \right)
$$

\n
$$
= 4E_{s} \sin^{2} \left(\frac{\pi}{M}(m-n) \right)
$$

\n
$$
d_{\min}^{2} = 4E_{s} \sin^{2} \left(\frac{\pi}{M} \right)
$$

$$
Q\left(\sqrt{\frac{2E_s \sin^2\left(\frac{\pi}{M}\right)}{N_0}}\right) \le P(\mathcal{E}) \le 2Q\left(\sqrt{\frac{2E_s \sin^2\left(\frac{\pi}{M}\right)}{N_0}}\right)
$$

Performance Bounds: MPSK

Performance Bounds: MPSK

 $E_b =$ *E^s* $\log_2(M)$

Performance (exact): M-PAM

$$
P(\mathcal{E}|\mathcal{H}_m) = 2Q\left(\sqrt{\frac{d^2}{2N_0}}\right) \qquad \text{(interior points)}
$$

2 "edge points"	
2 "edge points"	
1	
1	
1	
4	
4	
4	
4	
12E _s	
d = $\sqrt{\frac{12E_s}{M^2 - 1}}$	

M-2 "interior points"

$$
P(\mathcal{E}) = \frac{2}{M} \mathcal{Q} \left(\sqrt{\frac{d^2}{2N_0}} \right) + \frac{(M-2)}{M} 2 \mathcal{Q} \left(\sqrt{\frac{d^2}{2N_0}} \right)
$$

$$
=\frac{2(M-1)}{M}{\rm Q}\left(\sqrt{\frac{d^2}{2N_0}}\right)
$$

$$
P(\mathcal{E}) = \frac{2(M-1)}{M} \mathcal{Q}\left(\sqrt{\left[\frac{3}{M^2-1}\right] \frac{2E_s}{N_0}}\right)
$$

Performance (exact): M-QAM

\n
$$
M = M_p^2 = 4^i
$$
\n
$$
P(\mathcal{E}) = 1 - P(\mathcal{C})
$$
\n
$$
P(\mathcal{E}) = [P_{M_p - \text{PAM}}(\mathcal{C})]^2
$$
\n
$$
= \left[1 - \frac{2(M_p - 1)}{M_p} Q\left(\sqrt{\frac{d^2}{2N_0}}\right)\right]^2
$$
\n
$$
\approx \frac{4(\sqrt{M} - 1)}{\sqrt{M}} Q\left(\sqrt{\frac{3}{2(M - 1)}}\frac{2E_s}{N_0}\right)^2
$$
\n
$$
\approx \frac{4(\sqrt{M} - 1)}{\sqrt{M}} Q\left(\sqrt{\frac{3}{2(M - 1)}}\frac{2E_s}{N_0}\right)^2
$$
\n
$$
\approx \frac{4(\sqrt{M} - 1)}{\sqrt{M}} Q\left(\sqrt{\frac{3}{2(M - 1)}}\frac{2E_s}{N_0}\right)
$$

Performance (exact): M-QAM

Performance (exact): M-QAM

QAM/PSK Performance Comparison

QAM/PSK Performance Comparison

QPSK vs. BPSK Comparison 1 *Q*

QPSK

$$
P_s = 1 - \left[1 - Q\left(\sqrt{\frac{2E_b}{N_0}}\right)\right]^2 \approx 2Q\left(\sqrt{\frac{2E_b}{N_0}}\right)
$$

$$
P_b \approx Q\left(\sqrt{\frac{2E_b}{N_0}}\right)
$$
 (Gray mapping)

BPSK & Graylabeled QPSK

• This is actually exact when using Gray-mapped 4QAM/QPSK *is* actu r2*E^b N*⁰

Noise is independent on the In-phase and Quadrature dimensions

$$
\begin{aligned}\n\overbrace{\left(\begin{array}{c}\n\langle s_m, s_n\rangle = \Re\left\{\langle \bar{s}_m, \bar{s}_n\rangle\right\} = E_s\delta[m-n] \\
\phi_i(t) = \frac{s_i(t)}{\sqrt{E_s}}\n\end{array}\right)}^{(s_m, s_n) = E_s\delta[m-n]} \quad \|z - s_m\|^2 = \|z\|^2 + \|s_m\|^2 - 2z^t s_m \\
\overbrace{\left(\begin{array}{c}\n\langle s_m, s_n\rangle = \Re\left\{\langle \bar{s}_m, \bar{s}_n\rangle\right\} = E_s - 2z^t s_m \\
\overbrace{\left(\begin{array}{c}\n\bar{s}_m^t = \sqrt{E_s}\left(0 & 0 & \dots & 0 & 1 & 0 \dots & 0\right)\n\end{array}\right)}^{(1)}\right)}^{(1)} \\
\overbrace{\left(\begin{array}{c}\n\langle s_m, s_n\rangle = \Re\left\{\langle \bar{s}_m, \bar{s}_n\rangle\right\} = E_s - 2z^t s_m \\
\overbrace{\left(\begin{array}{c}\n\bar{s}_m^t = \sqrt{E_s}\left(0 & 0 & \dots & 0 & 1 & 0 \dots & 0\right)\n\end{array}\right)}^{(1)}\right)}^{(1)} \\
\overbrace{\left(\begin{array}{c}\n\bar{s}_m^t = \sqrt{E_s}\left(0 & 0 & \dots & 0 & 1 & 0 \dots & 0\right)\n\end{array}\right)}^{(1)}\n\end{aligned}
$$

$$
f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_0) = \mathcal{N}_M\left(\mathbf{z}; \mathbf{s}_0; \frac{N_0}{2}\mathbf{I}\right)
$$

= $\mathcal{N}_1(z_0; \sqrt{E_s}; N_0/2) \prod_{i=1}^{M-1} \mathcal{N}_1(z_i; 0; N_0/2)$

Given \mathcal{H}_0 , $\{z_0(u), z_1(u), \ldots z_{M-1}(u)\}$ are mutually independent

 $P(\mathcal{C}|\mathcal{H}_0, z_0(u) = z_0) = \Pr\{z_1(u) < z_0, z_2(u) < z_0, \ldots z_{M-1}(u) < z_0 | \mathcal{H}_0, z_0(u) = z_0\}$

$$
= \prod_{i=1}^{M-1} \Pr\left\{z_i(u) < z_0 | \mathcal{H}_0, z_0(u) = z_0\right\}
$$

$$
=\prod_{i=1}^{M-1}\operatorname{Pr}\left\{z_i(u)
$$

$$
= \prod_{i=1}^{M-1} \left[1 - \mathcal{Q}\left(\sqrt{\frac{2}{N_0}} z_0\right) \right]
$$

$$
= \left[1 - \mathcal{Q}\left(\sqrt{\frac{2}{N_0}}z_0\right)\right]^{M-1}
$$

$$
P(\mathcal{C}|\mathcal{H}_0) = \int_{-\infty}^{\infty} P(\mathcal{C}|\mathcal{H}_0, z_0(u) = z_0) f_{z_0(u)}(z_0|\mathcal{H}_0) dz_0
$$

$$
= \int_{-\infty}^{\infty} \left[1 - Q\left(\sqrt{\frac{2}{N_0}}z_0\right)\right]^{M-1} \mathcal{N}_1(z_0; \sqrt{E_s}; N_0/2) dz_0
$$

$$
= \int_{-\infty}^{\infty} \left[\int_{-\infty}^{x+\sqrt{\frac{2E_b}{N_0}\log_2(M)}} \mathcal{N}_1(v;0;1) dv \right]^{M-1} \mathcal{N}_1(x;0;1) dx
$$

$$
= \int_{-\infty}^{\infty} \left[F\left(x + \sqrt{\frac{2E_b}{N_0} \log_2(M)}\right) \right]^{M-1} \mathcal{N}_1(x; 0; 1) dx
$$

 $F(x) = \text{cdf of a standard Gaussian}$

$$
P(\mathcal{E}) = P(E|\mathcal{H}_0) = 1 - \int_{-\infty}^{\infty} \left[F\left(x + \sqrt{\frac{2E_b}{N_0} \log_2(M)}\right) \right]^{M-1} \mathcal{N}_1(x; 0; 1) dx
$$

$$
P(\mathcal{E}) = P(E|\mathcal{H}_0) = 1 - \int_{-\infty}^{\infty} \left[F\left(x + \sqrt{\frac{2E_b}{N_0} \log_2(M)}\right) \right]^{M-1} \mathcal{N}_1(x; 0; 1) dx
$$

performance improves as M increases — opposite of QASK

performance improves as M increases — opposite of QASK

How far can we take this?

$$
\lim_{M \to \infty} P_M(\mathcal{E}) = 1 - \lim_{M \to \infty} P_M(\mathcal{C})
$$
\n
$$
= 1 - \exp\left[\lim_{M \to \infty} \ln\left(P_M(\mathcal{C})\right)\right]
$$
\n
$$
\lim_{M \to \infty} \ln\left(P_M(\mathcal{C})\right) = \begin{cases}\n-\infty & \frac{E_b}{N_0} < \ln(2) \\
0 & \frac{E_b}{N_0} \ge \ln(2)\n\end{cases}
$$

(use L'Hopital's Rule) *P (use L'Hopital's R N*⁰ **l**e)

M-ary Orthogonal (coherent demod) ln (*PM*(*C*)) = (1 *^E^b <* ln(2)

How far can we take this?

$$
\lim_{M \to \infty} P_M(\mathcal{E}) = \begin{cases} 1 & \frac{E_b}{N_0} < \ln(2) \\ 0 & \frac{E_b}{N_0} \ge \ln(2) \end{cases}
$$
\n
$$
\left(\frac{1}{2} \right)^2 & \frac{E_b}{N_0} < \ln(2)
$$

$$
\lim_{M \to \infty} P_b(M) = \begin{cases} 1/2 & \frac{E_b}{N_0} < \ln(2) \\ 0 & \frac{E_b}{N_0} \ge \ln(2) \end{cases}
$$

$$
\eta_{\text{bits/dim}} = \frac{\log_2(M)}{M}
$$

lim $M \rightarrow \infty$ $\eta_{\rm bits/dim} = 0$

ln(2) is a threshold on Eb/No for which perfect communication occurs

spectral efficiency (bps/Hz) goes to zero as M increases

We will see that this result shows that orthogonal modulation achieves Shannon Capacity for the AWGN as spectral efficiency goes to 0

Also, at finite spectral efficiency, the capacity results will show that similar threshold results hold, but for larger values of Eb/No

Eb/No $= -1.6$ dB is the smallest value of Eb/No for reliable communications on the AWGN channel

- Other "orthogonal-like" signal sets exhibit similar large M performance trends
	- Bi-orthogonal, simplex
- Result also occurs with phase non-coherent detection

AWGN Capacity (preview)

 $Eb/No = -1.6 dB$ is the smallest value of Eb/No for reliable communications on the AWGN channel

Modulation Comparison to Capacity

About 10 dB Eb/No gap to capacity for these uncoded QAM

Modulation Comparison to Capacity

About 10 dB Eb/No gap to capacity for orthogonal too

Detection/Demod Topics

- Maximum A Posteriori decision rule for vector-AWGN channel
- Exact performance for binary modulations
- Minimum distance decision rule for M-ary modulation over AWGN
- Performance bounds
	- *• Performance of common M-ary modulations*
- Continuous time model
	- Likelihood functional, sufficient statistics
- Average and generalized likelihood
	- Phase non-coherent demodulation
	- Soft-out demodulation

$$
\mathcal{H}_m: \quad \mathbf{z} = \mathbf{s}_m + \mathbf{w} \qquad D \times 1
$$
\n
$$
f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m) \equiv \exp\left(\frac{-\|\mathbf{z} - \mathbf{s}_m\|^2}{N_0}\right)
$$
\n
$$
\equiv e^{-E_m/N_0} \exp\left(\frac{2\mathbf{s}_m^t \mathbf{z}}{N_0}\right)
$$

vector observation — finite number of random variables

$$
\mathcal{H}_m: \quad r(u,t) = s_m(t) + n(u,t) \quad t \in \mathcal{T}
$$

 $\mathcal{H}_m:$ $r(u,t) = s_m(t) + n(u,t)$ $t \in \mathcal{T}$ waveform observation — waveform observation — uncountably infinite random variables

$$
L(r|\mathcal{H}_m) \equiv \exp\left(\frac{-1}{N_0} \left[\|\mathbf{s}_m\|^2 - 2\langle \mathbf{r}, \mathbf{s}_m \rangle \right] \right)
$$

$$
= \exp\left(\frac{-1}{N_0} \left[\int_{\mathcal{T}} s_m^2(t)dt - 2 \int_{\mathcal{T}} r(t)s_m(t)dt \right] \right)
$$

$$
= e^{-E_m/N_0} \exp\left(\frac{2}{N_0} \int_{\mathcal{T}} r(t)s_m(t)dt \right)
$$

AWGN Continuous time likelihood functional (replaces conditional pdf/pmf)

$$
r^{(W)}(t) = s_m(t) + n(u, t)
$$

\n
$$
S_n(f) = N_0/2
$$

\n
$$
R_n(\tau) = \frac{N_0}{2} \delta(\tau)
$$

\n
$$
R_n(w)(\tau) = N_0 W \operatorname{sinc}(2W\tau)
$$

\n
$$
T_{\text{sample}} = \frac{1}{2W}
$$

\n
$$
S_{n(w)}(f) = \frac{N_0}{2} \operatorname{rect}(f/(2W))
$$

\n
$$
R_{n(w)}(\tau) = N_0 W \operatorname{sinc}(2W\tau)
$$

$$
\mathcal{H}_m: \quad \mathbf{r}^{(W)}(u) = \mathbf{s}_m^{(W)} + \mathbf{n}^{(W)}(u) \qquad (2WT) \times 1
$$

$$
\mathbf{n}^{(W)}(u) \sim \mathcal{N}_{2WT}\left(\cdot ; \mathbf{0} ; N_0 W \mathbf{I}\right)
$$

$$
f_{\mathbf{r}^{(W)}(u)}(\mathbf{r}^{(W)}|\mathcal{H}_m) \equiv \exp\left(\frac{-\|\mathbf{r}^{(W)} - \mathbf{s}_m^{(W)}\|^2}{2WN_0}\right)
$$

$$
\equiv e^{-\|\mathbf{s}_m^{(W)}\|^2/(2WN_0)} \exp\left(\frac{2\left[\mathbf{s}_m^{(W)}\right]^\mathsf{t}\mathbf{r}^{(W)}}{2WN_0}\right)
$$

$$
\frac{1}{2WN_0} ||\mathbf{s}_m^{(W)}||^2 = \frac{1}{N_0} \sum_i \left(s_m^{(W)} \left(\frac{i}{2W} \right) \right)^2 \frac{1}{2W}
$$

$$
\frac{1}{2WN_0} \left[\mathbf{s}_m^{(W)} \right]^{\text{t}} \mathbf{r}^{(W)} = \frac{1}{N_0} \sum_i r^{(W)} \left(\frac{i}{2W} \right) s_m^{(W)} \left(\frac{i}{2W} \right) \frac{1}{2W}
$$

As W goes to infinity, this converges to the integral

More rigorous development from a generalized Fourier Series expansion of the observed signal - Karhunen-Loeve Expansion

$$
\mathcal{H}_m: \quad r(u,t) = s_m(t) + n(u,t) \quad t \in \mathcal{T}
$$

$$
\mathcal{H}_m: \quad R(u, i) = S_m(i) + N(u, i) \quad i = 1, 2, 3, ...
$$

$$
R(u,i) = \int_{\mathcal{T}} r(u,t)\phi_i(t)dt
$$

$$
S_m(i) = \int_{\mathcal{T}} s_m(t)\phi_i(t)dt
$$

$$
N(u,i) = \int_{\mathcal{T}} n(u,t)\phi_i(t)dt
$$

Continuous Time AWGN Likelihood Functional *R*(*k*) $\frac{1}{2}$

More rigorous development from a generalized Fourier Series riore rigorous development from a generalized rourier beries
expansion of the observed signal - Karhunen-Loeve Expansion

$$
N(u,i) = \int_{\mathcal{T}} n(u,t)\phi_i(t)dt
$$

$$
\mathbb{E}\left\{N(u,i)\right\} = \int_{\mathcal{T}} \mathbb{E}\left\{n(u,t)\right\} \phi_i(t)dt = 0
$$

$$
\mathbb{E}\left\{N(u,i)N(u,k)\right\} = \int_{\mathcal{T}} \int_{\mathcal{T}} \phi_i(t_1) \mathbb{E}\left\{n(u,t_1)n(u,t_2)\right\} \phi_k(t_2)dt_1dt_2
$$

$$
= \int_{\mathcal{T}} \int_{\mathcal{T}} \phi_i(t_1)K_n(t_1,t_2) \phi_k(t_2)dt_1dt_2
$$

Karhunen-Loeve Expansion comes from solving

$$
\int_{\mathcal{T}} K_n(t_1, t_2) \phi_k(t_2) dt_2 = \lambda_k \phi_k(t_1)
$$

This implies that the noise coefficients are uncorrelated

$$
\mathbb{E}\left\{N(u,i)N(u,k)\right\} = \lambda_k \delta[i-k]
$$

For the generalized FS we have

$$
\sum_i X(i)Y(i) = \int_{\mathcal{T}} x(t)y(t)dt
$$

In the limiting case of AWGN

z
Z \overline{I} $K_n(t_1, t_2) \phi_k(t_2) dt_2 =$ z
Z \overline{I} *N*⁰ 2 $\delta(t_1 - t_2)\phi_k(t_2)dt_2 =$ *N*⁰ $\frac{1}{2}$ $\phi_k(t_1)$ Any CONS works

Summary of Karhunen-Loeve Expansion for AWGN Limiting Sase $\frac{1}{2}$ $\frac{1}{2}$ $\frac{1}{2}$ *T Sm*(*t*)*i*(*t*)*dt*

$$
N(u, i) = \int_{\mathcal{T}} n(u, t)\phi_i(t)dt
$$
 Any CONS

$$
\mathbb{E}\{N(u, i)N(u, k)\} = 0
$$

$$
\mathbb{E}\{N(u, i)N(u, k)\} = \frac{N_0}{2}\delta[i - k]
$$

Jointly Gaussian (iid) coefficients

Related Facts About Correlating AWGN

$$
N_a(u) = \int_{\mathcal{T}} a(t)n(u, t)dt
$$

$$
N_b(u) = \int_{\mathcal{T}} b(t)n(u, t)dt
$$

$$
\mathbb{E}\{N_a(u)\} = \mathbb{E}\{N_b(u)\} = 0
$$

$$
\mathbb{E}\{N_a^2(u)\} = \frac{N_0}{2} \int_{\mathcal{T}} a^2(t)dt
$$

$$
\mathbb{E}\{N_b^2(u)\} = \frac{N_0}{2} \int_{\mathcal{T}} b^2(t)dt
$$

$$
\mathbb{E}\{N_a(u)N_b(u)\} = \frac{N_0}{2} \int_{\mathcal{T}} a(t)b(t)dt
$$

KL expansion for AWGN leads to Likelihood Functional

$$
f(\{R(i)\}_{i=1}^{N}|\mathcal{H}_{m}) \equiv \exp\left(\frac{-1}{N_{0}}\sum_{i=1}^{N}S_{m}^{2}(i) + \frac{2}{N_{0}}\sum_{i=1}^{N}R(i)S_{m}(i)\right)
$$

$$
\lim_{N \to \infty} f(\{R(i)\}_{i=1}^{N}|\mathcal{H}_{m}) \equiv \exp\left(\frac{-1}{N_{0}}\int_{\mathcal{T}}s_{m}^{2}(t)dt + \frac{2}{N_{0}}\int_{\mathcal{T}}r(t)s_{m}(t)dt\right)
$$

CONS = { $\phi_1(t), \phi_2(t), \ldots, \phi_D(t)$ } $\bigcup \{\phi_{D+1}(t), \phi_{D+2}(t), \phi_{D+3}(t), \ldots\}$

Orthonormal basis for signal space

orthonormal completion of first D functions

KL expansion for AWGN leads to Likelihood Functional

$$
f(\lbrace R(i)\rbrace_{i=1}^{D+k} | \mathcal{H}_m) = \mathcal{N}_{D+k} \left(\begin{bmatrix} R(1) \\ R(2) \\ \vdots \\ R(D) \\ \vdots \\ R(D+1) \\ \vdots \\ R(D+k) \end{bmatrix}; \begin{bmatrix} S_m(1) \\ S_m(2) \\ \vdots \\ S_m(D) \\ \vdots \\ 0 \end{bmatrix}; \begin{bmatrix} \frac{N_0}{2} \mathbf{I}_D & \mathbf{O} \\ \mathbf{O} & \frac{N_0}{2} \mathbf{I}_k \end{bmatrix} \right)
$$

$$
= \mathcal{N}_D \left(\begin{bmatrix} R(1) \\ R(2) \\ \vdots \\ R(2) \\ \vdots \\ R(D) \end{bmatrix}; \begin{bmatrix} S_m(1) \\ S_m(2) \\ \vdots \\ S_m(D) \end{bmatrix}; \frac{N_0}{2} \mathbf{I}_D \right) \mathcal{N}_k \left(\begin{bmatrix} R(D+1) \\ R(D+2) \\ \vdots \\ R(D+k) \end{bmatrix}; \mathbf{0}; \frac{N_0}{2} \mathbf{I}_k \right)
$$

Continuous Time AWGN Likelihood Functional A' 6 V $\mathbf V$ 6 R . $\overline{}$ **1** \blacktriangleright 7 $\ddot{\mathbf{S}}$. $\overline{?}$ \mathbf{r} \mathbf{r} 7 *N*⁰ **20** I*D Nk* $\overline{}$ E
B
B 6 $\overline{1}$ \overline{a} $\overline{}$ *R*(*D* + 2) . \overline{a} \mathbf{A} 1al 7

KL expansion for AWGN leads to Likelihood Functional $\overline{1}$ SN le .
ما جام *د Sm*(*D*) 1
1000 - J. Europiano

$$
f(\lbrace R(i)\rbrace_{i=1}^{D+k}|\mathcal{H}_m) \equiv \mathcal{N}_D\left(\begin{bmatrix}R(1)\\R(2)\\ \vdots\\R(D)\end{bmatrix};\begin{bmatrix}S_m(1)\\S_m(2)\\ \vdots\\S_m(D)\end{bmatrix};\frac{N_0}{2}\mathbf{I}_D\right)
$$

$$
\equiv \exp\left(\frac{-\Vert\mathbf{z}-\mathbf{s}_m\Vert^2}{N_0}\right)
$$

$$
\mathbf{z} = \begin{bmatrix}R(1) & R(2) & \cdots & R(D)\end{bmatrix}^{\mathrm{t}}
$$

$$
\mathbf{s}_m = \begin{bmatrix}S_m(1) & S_m(2) & \cdots & S_m(D)\end{bmatrix}^{\mathrm{t}}
$$

Vector-AWGN Model is Equivalent to Processing Continuous AWGN observation optimally!

$$
\mathbf{z}^{\mathbf{t}}\mathbf{s}_{m} = \int_{\mathcal{T}} r(t)s_{m}(t)dt = \sum_{i=1}^{D} R(i)S_{m}(i)
$$

$$
E_{m} = \int_{\mathcal{T}} s_{m}^{2}(t)dt = \sum_{i=1}^{D} S_{m}^{2}(i)
$$

$$
L(r|\mathcal{H}_m) \equiv \exp\left(\frac{-1}{N_0} \left[\|\mathbf{s}_m\|^2 - 2\langle \mathbf{r}, \mathbf{s}_m \rangle \right] \right)
$$

$$
= \exp\left(\frac{-1}{N_0} \left[\int_{\mathcal{T}} s_m^2(t)dt - 2 \int_{\mathcal{T}} r(t)s_m(t)dt \right] \right)
$$

$$
= e^{-E_m/N_0} \exp\left(\frac{2}{N_0} \int_{\mathcal{T}} r(t)s_m(t)dt \right)
$$

$$
= e^{-E_m/N_0} \exp\left(\frac{2}{N_0} \mathbf{z}^{\mathbf{t}} \mathbf{s}_m \right)
$$

Vector-AWGN Model is Equivalent to Processing Continuous AWGN observation optimally!

(correlation to orthonormal basis)

Vector-AWGN model is a model for the output of the bank of correlates to the orthonormal basis for the signal space

Sufficient Stats and Related Topics

- Note that the processing from $r(t)$ to **z** is is not reversible
	- Cannot recover r(t) from **^z**
- In general, how do we know that we are not throwing out useful information?
	- Notion of a *set of sufficient statistics*
	- Engineering lingo (Wozencraft & Jacobs)
		- *• Theorem of Reversibility*
		- *• Theorem of Irrelevance*

Theorem of Reversibility

• Any reversible (invertible) signal processing operation can be performed on the observation without losing information relevant to the decision problem

- 1. A nonzero centroid does not help performance and wastes energy
- 2. A unitary transformation of the signals (e.g., rotation, reflection) does not affect performance in AWGN

Theorem of Reversibility

Colored Noise MAP receiver realized using a whitening filter

Theorem of Irrelevance

Suppose we have two observations

$$
f(\mathbf{z}_1, \mathbf{z}_2 | \mathcal{H}_m) = f(\mathbf{z}_1 | \mathbf{z}_2, \mathcal{H}_m) f(\mathbf{z}_2 | \mathcal{H}_m)
$$

If the following holds

$$
f(\mathbf{z}_1|\mathbf{z}_2,\mathcal{H}_m)=f(\mathbf{z}_1|\mathbf{z}_2) \quad m=0,1,\ldots M-1
$$

Then we say that z1 is irrelevant given z2 for the purposes of making a decision on the hypotheses

We have used this when dropping multiplicative terms and in discarding the AWGN outside the signal space

Set of Sufficient Statistics

A set of sufficient statistics for a hypothesis testing problem is a function of the observation that makes the observation irrelevant

$$
f(\mathbf{z}, \mathbf{g}(\mathbf{z}) | \mathcal{H}_m) = f(\mathbf{z} | \mathbf{g}(\mathbf{z}), \mathcal{H}_m) f(\mathbf{g}(\mathbf{z}) | \mathcal{H}_m)
$$

= $f(\mathbf{z} | \mathbf{g}(\mathbf{z})) f(\mathbf{g}(\mathbf{z}) | \mathcal{H}_m)$ $m = 0, 1, ... M - 1$
 $\equiv f(\mathbf{g}(\mathbf{z}) | \mathcal{H}_m)$

Examples:

 ${\{\mathbf{z}^{\text{t}}\mathbf{s}_m\}}_{m=0}^{M-1}$ is a set of sufficient stats for the vector AWGN channel

 \bigcap \overline{I} $r(t)s_m(t)dt\bigg\}^{M-1}$ $m=0$ is a set of sufficient stats for the AWGN channel

$$
\left\{\int_{\mathcal{T}} r(t) \phi_i(t) dt \right\}_{i=0}^D \text{ is a set of sufficient stats for the AWGN channel}
$$

Set of Sufficient Statistics

If you start with the likelihood (functional) and you simplify to the only hypothesis-dependent terms that are a function of the observation, then these are a set of sufficient stats

Whenever you have a set of sufficient statistics, they can be treated as an equivalent observation and the hypothesis testing problem can be reformulated using the equivalent observation

Example: We reformulated the waveform AWGN channel problem in terms of the equivalent vector model which we now see is a set of sufficient stats

Why not use the other set of sufficient stats from previous slide?

(correlation to signals)

Requires M >= D correlators

Post-correlator model for processing of previous slide

$$
\mathcal{H}_m: \quad \mathbf{r}(u) = \mathbf{v}_m + \mathbf{n}(u)
$$

$$
\mathbf{v}_m = \left[\begin{array}{cccc} \langle \mathbf{s}_m, \mathbf{s}_0 \rangle & \langle \mathbf{s}_m, \mathbf{s}_1 \rangle & \langle \mathbf{s}_m, \mathbf{s}_2 \rangle & \cdots & \langle \mathbf{s}_m, \mathbf{s}_{M-1} \rangle \end{array} \right]^{\mathrm{t}}
$$

 $m_n = 0$

$$
\mathbf{K_n} = \frac{N_0}{2}\left[\begin{array}{cccc} \langle \pmb{s}_0, \pmb{s}_0\rangle & \langle \pmb{s}_0, \pmb{s}_1\rangle & \langle \pmb{s}_0, \pmb{s}_2\rangle & \ldots & \langle \pmb{s}_0, \pmb{s}_{M-1}\rangle \\[1mm] \langle \pmb{s}_1, \pmb{s}_0\rangle & \langle \pmb{s}_1, \pmb{s}_1\rangle & \langle \pmb{s}_1, \pmb{s}_2\rangle & \ldots & \langle \pmb{s}_1, \pmb{s}_{M-1}\rangle \\[1mm] \langle \pmb{s}_1, \pmb{s}_0\rangle & \langle \pmb{s}_1, \pmb{s}_1\rangle & \langle \pmb{s}_1, \pmb{s}_2\rangle & \ldots & \langle \pmb{s}_1, \pmb{s}_{M-1}\rangle \\[1mm] \vdots & \vdots & \ddots & \vdots \\[1mm] \langle \pmb{s}_{M-1}, \pmb{s}_0\rangle & \langle \pmb{s}_{M-1}, \pmb{s}_1\rangle & \langle \pmb{s}_{M-1}, \pmb{s}_2\rangle & \ldots & \langle \pmb{s}_{M-1}, \pmb{s}_{M-1}\rangle \end{array}\right]
$$

This matrix of inner products is called the Gramian of the signal set

(correlation to signals)

- This processing is not preferred because
	- More correlates than needs (high complexity)
	- Noise vector covariance matrix will have rank D which means it is singular unless D=M
- For orthogonal signaling, the two approaches are the same!

This post-correlator model illustrates that the performance in AWGN is completely determined by the Gramian of the signal set — i.e., the inner products between signals

Complex BB CT Likelihood Functional

Recall:

$$
\mathcal{H}_m: \quad r(u,t) = s_m(t) + n(u,t) \quad t \in \mathcal{T} \quad \text{(narrowband)}
$$

$$
\mathcal{H}_m: \quad \bar{r}(u,t) = \bar{s}_m(t) + \bar{n}(u,t) \qquad t \in \mathcal{T} \qquad \qquad \text{(complex BB)}
$$

H^m : *r*(*u, t*) = *sm*(*t*) + *n*(*u, t*) *t* 2 *T* I and Q components of complex BB equivalent AWGN are each AWGN processes that are independent

$$
\langle \boldsymbol{r}, \boldsymbol{s}_m \rangle = \int_{\mathcal{T}} r(t) s_m(t) dt = \Re \{ \langle \bar{\boldsymbol{r}}, \bar{\boldsymbol{s}}_m \rangle \} = \Re \{ \int_{\mathcal{T}} r(t) s_m^*(t) dt \}
$$
Complex BB CT Likelihood Functional

$$
L(r|\mathcal{H}_m) \equiv \exp\left(\frac{-1}{N_0} \left[\lVert s_m \rVert^2 - 2\langle r, s_m \rangle \right] \right)
$$

\n
$$
= \exp\left(\frac{-1}{N_0} \left[\int_{\mathcal{T}} s_m^2(t)dt - 2 \int_{\mathcal{T}} r(t)s_m(t)dt \right] \right)
$$

\n
$$
= e^{-E_m/N_0} \exp\left(\frac{2}{N_0} \int_{\mathcal{T}} r(t)s_m(t)dt \right)
$$

\n
$$
= \exp\left(\frac{-1}{N_0} \left[\int_{\mathcal{T}} |\bar{s}_m(t)|^2 dt - 2\Re\left\{ \int_{\mathcal{T}} r(t)s_m^*(t)dt \right\} \right] \right)
$$

\n
$$
= L(\bar{r}|\mathcal{H}_m)
$$

\n(complex BB)

Detection of a Digital Sequence

For the PSD, we considered a sequence of digital symbols sent through the channel

linear (QASK) modulation

$$
\overline{x}(u,t) = \sum_{k} \overline{X}_{k}(u)p(t - kT)
$$

 $\overline{X}_k(u)$ ~ independent, distributed over QASK constellation

- So far we have considered only the "one shot" detection problem
	- Let's use the continuous time likelihood functional to solve this sequence detection problem

Detection of a Digital Sequence

memoryless (nonlinear) modulation

$$
\overline{x}(u,t) = \sum_{k} \overline{s}_{X_k(u)}(t - kT)
$$

$$
X_k(u) \in \{0, 1, \dots M - 1\} \quad \text{(independent)}
$$

$$
\overline{s}_m(t) \text{(lasts} \le T \text{ seconds)}
$$

$$
L(\bar{\mathbf{r}}|\{X_k(u) = a_k\}_k) = \exp\left(\frac{-1}{N_0} \left[\int_{\mathcal{T}} |\bar{x}(u, t; \mathbf{a})|^2 dt - 2\Re \left\{ \int_{\mathcal{T}} \bar{r}(t) \bar{x}^*(u, t; \mathbf{a}) dt \right\} \right] \right)
$$

$$
= \exp\left(\frac{-1}{N_0} \left[\sum_k \int_{kT}^{(k+1)T} |\bar{s}_{a_k}(t)|^2 dt - 2\Re \left\{ \sum_k \int_{kT}^{(k+1)T} \bar{r}(t) \bar{s}_{a_k}^*(t) dt \right\} \right] \right)
$$

$$
= \prod_k \exp\left(\frac{-1}{N_0} \left[\int_{kT}^{(k+1)T} |\bar{s}_{a_k}(t)|^2 dt - 2\Re \left\{ \int_{kT}^{(k+1)T} \bar{r}(t) \bar{s}_{a_k}^*(t) dt \right\} \right] \right)
$$

$$
= \prod_k L(\bar{\mathbf{r}}_k | X_k(u) = a_k)
$$

For independent modulation symbols, the likelihood functional factors and the optimal processing is to repeat the one-shot MAP detector each symbol time

MAP Receiver in AWGN

(correlation to orthonormal basis)

Can think of this as just resetting the one-shot detector and repeating each symbol time

Example: MFSK Orthogonal

Correlation vs Matched Filter

$$
x(t) * v(t) = \int x(\tau)v(t - \tau)d\tau
$$

$$
x(t) * v^*(-t) = \int x(\tau)v^*(\tau - t)d\tau
$$

$$
x(t) * v^*(-t)|_{t=kT} = \int x(\tau)v^*(\tau - kT)d\tau
$$

Correlation vs Matched Filter

Correlation vs Matched Filter

bank of correlates for signal that last 3T

matched filter

single matched-filter required even if signal phi lasts multiple symbol

MAP Receiver in AWGN

(matched-filters to orthonormal basis)

correlator form is common with rect-pulses and called an "integrate and dump"

linear modulation

(consider arbitrary pulse shape)

$$
\overline{x}(u,t) = \sum_{k} \overline{X}_{k}(u)p(t - kT)
$$

 $\overline{X}_k(u) \sim$ independent, distributed over QASK constellation

$$
L(\bar{r}|\{\bar{X}_k(u) = \bar{a}_k\}_k) = \exp\left(\frac{-1}{N_0} \left[\int_{\mathcal{T}} |\bar{x}(u, t; \bar{\mathbf{a}})|^2 dt - 2\Re\left\{ \int_{\mathcal{T}} \bar{r}(t)\bar{x}^*(u, t; \bar{\mathbf{a}}) dt \right\} \right] \right)
$$

$$
= \exp\left(\frac{-1}{N_0} \left[\int_{\mathcal{T}} |\bar{x}(u, t; \bar{\mathbf{a}})|^2 dt \right] \right) \exp\left(\frac{-1}{N_0} \left[2\Re\left\{ \int_{\mathcal{T}} \bar{r}(t) \sum_k \bar{a}_k^* p^*(t - kT) dt \right\} \right] \right)
$$

$$
= \exp\left(\frac{-1}{N_0} \left[\int_{\mathcal{T}} |\bar{x}(u, t; \bar{\mathbf{a}})|^2 dt \right] \right) \exp\left(\frac{-1}{N_0} \left[2\Re\left\{ \sum_k \bar{a}_k^* \int_{\mathcal{T}} \bar{r}(t) p^*(t - kT) dt \right\} \right] \right)
$$

$$
\bar{z}_k = \int \bar{r}(t) p^*(t - kT) dt
$$

 $\{\bar{z}_k\}_k$ is a set of sufficient statistics

complex baseband matched filter outputs

complex baseband

narrowband signal processing for real pulse

$$
\overline{r}(t) \longrightarrow p^*(-t) \qquad \overline{s}_p(t; \overline{\mathbf{a}}) + \overline{n}_p(u, t) \qquad \overline{z}_k \qquad \overline{z}_k
$$

$$
\overline{s}_p(t; \overline{\mathbf{a}}) = \left[\sum_i \overline{a}_i p(t - iT)\right] * p^*(-t)
$$

$$
= \left[\sum_i \overline{a}_i \delta(t - iT)\right] * p(t) * p^*(-t)
$$

$$
= \left[\sum_i \overline{a}_i \delta(t - iT)\right] * R_p(t)
$$

 $\bar{a}_i R_p(t-iT)$

 $=$ \sum

i

$$
\bar{n}_p(u,t) = \bar{n}(u,t) * p^*(-t)
$$

$$
R_{\bar{n}_p}(\tau) = N_0 R_p(\tau)
$$

$$
\bar{z}_k(u) = \left[\bar{s}_p(t; \bar{\mathbf{a}}) + \bar{n}_p(u, t)\right]|_{t=k}
$$

$$
= \sum_i \bar{a}_i R_p((k-i)T) + \bar{n}_k(u)
$$

 $\mathbb{E} \left\{ \bar{n}_{k+m}(u)\bar{n}_{k}^{*}(u) \right\} = N_{0}R_{p}(mT)$

$$
R_p(mT) = p(t) * p^*(-t)|_{t=mT} = C\delta[m]
$$

$$
\bar{z}_k(u) = \bar{a}_k + \bar{w}_k(u)
$$

$$
\mathbb{E}\left\{\bar{w}_{k+m}(u)\bar{w}_k^*(u)\right\} = N_0\delta[m]
$$

Nyquist Condition on pulse shape When p(t) satisfies the Nyquist condition

- There is no inter symbol interference
- The noise at the output of the MF is CC-AWGN

The Nyquist condition is satisfied for any pulse that is zero outside of [0,T]

Can a pulse that lasts longer than T satisfy this?

Nyquist Condition for No ISI

 $time$ domain

$$
R_p(mT) = p(t) * p^*(-t)|_{t=mT} = C\delta[m]
$$

$$
\mathbb{FT}\left\{R_p(t)\right\} = |P(f)|^2
$$

frequency domain

$$
\frac{1}{T}\sum_{k}|P(f-k/T)|^2 = C
$$

folded-spectrum should be flat

Nyquist Condition on pulse shape (freq domain)

Nyquist Pulse Shape: sinc()

 $|P(f)|^2 = T \text{rect}(fT)$

 $R_p(t) = \text{sinc}(t/T)$

 $P(f) = \sqrt{T} \text{rect}(fT)$

$$
p(t) = \frac{1}{\sqrt{T}}\mathrm{sinc}(t/T)
$$

Nyquist Condition on pulse shape (freq domain)

Nyquist Pulse Shape: sinc()

sample waveform for 4PAM with sinc() pulse shape (matched filter output)

Nyquist Pulse Shape: Raised Cosine Spectrum

$$
|P(f)|^2 = \begin{cases} T & |f| < \frac{1-\beta}{2T} \\ \frac{T}{2} \left[1 - \sin\left(\frac{\pi T}{\beta} \left(f - \frac{1}{2T}\right)\right) \right] & \frac{1-\beta}{2T} \le |f| \le \frac{1+\beta}{2T} \\ 0 & |f| > \frac{1+\beta}{2T} \end{cases}
$$

1 42(*t/T*)² $\beta \in [0, 1)$ fractional excess bandwidth

Nyquist Pulse Shape: Raised Cosine Pulse Rp

note that pulse correlation passes through zero at integer multiples of T

Nyquist Pulse Shape: Raised Cosine Spectrum ⁰ *[|]f[|] >* 1+ 2*T*

 $R_p(t) = \text{sinc}(t/T) \frac{\text{cos}(\beta \pi t/T)}{4 \beta^2 (t/T)}$ $1 - 4\beta^2(t/T)^2$ "raised cosine pulse"

$$
P(f) = |P(f)|
$$

 $p(t)=4\beta \frac{\cos((1+\beta)\pi t/T)+\sin((1-\beta)\pi(t/T)) \left[4\beta(t/T)\right]^{-1}}{\sqrt{\pi}}$ π $\overline{}$ $T\,[1-(4\beta t/T)^2]$ "root raised cosine pulse"

These are built into Matlab!

rcosdesign(0.35,40,N_sps,'norm'); Raised cosine

rcosdesign(0.35,40,N_sps,'sqrt'); Roots-Raised cosine

beta = 0.35, truncated to 40 symbols length, number of samples per symbol

Nyquist Pulse Shape: Root Raised Cosine Pulse

note that pulse does not pass through zero at integer multiples of T

Nyquist Pulse Shape: Raised Cosine Spectrum

Nyquist Pulse Shape: Raised Cosine Spectrum

Signal trajectory in the I/Q plane with RRC pulse shaping

PSK with RRC has envelope variation

 $I(t)$

QASK Modulation with Nyquist Pulse Shaping

Detection/Demod Topics

- Maximum A Posteriori decision rule for vector-AWGN channel
- Exact performance for binary modulations
- Minimum distance decision rule for M-ary modulation over AWGN
- Performance bounds
	- *• Performance of common M-ary modulations*
- Continuous time model
	- Likelihood functional, sufficient statistics
- Average and generalized likelihood
	- Phase non-coherent demodulation
	- Soft-out demodulation

Composite Hypothesis Testing

- The observation model is a function of a parameter or a set of parameters
	- Nuisance parameters
- If we have a statistical model for the nuisance parameters
	- Average them out this is called *average likelihood*
		- Same as original likelihood, just a two step process
- If no statistical model is assumed
	- Can maximize over the parameters along with the hypothesis
		- Called *generalized likelihood* (joint likelihood)
			- Ad hoc in general

Composite Hypothesis Testing - Topics

- Basic concepts and definitions
- Phase noncoherent detection
- Differential encoding of PSK and differentially coherent detection
- Soft-output demappers
	- Get soft decisions out of the M-ary decision device
		- transition to coding

Composite Hypothesis Testing

Average Likelihood

$$
f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m) = \int f_{\mathbf{z}(u)|\Theta(u)}(\mathbf{z}|\theta, \mathcal{H}_m) f_{\Theta(u)}(\theta|\mathcal{H}_m) d\theta
$$

$$
= \int f_{\mathbf{z}(u)|\Theta(u)}(\mathbf{z}|\theta, \mathcal{H}_m) f_{\Theta(u)}(\theta) d\theta \quad (\Theta(u) \text{ independent of hypothesis})
$$

Generalized Likelihood

$$
g_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m) = \max_{\theta} f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m; \theta)
$$

$$
= f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m; \hat{\theta}_m)
$$

$$
\hat{\theta}_m = \arg \max_{\theta} f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m; \theta)
$$

Phase Noncoherent Detection

Use average likelihood with nuisance parameter being the incoming carrier phase

$$
f_{\mathbf{z}(u)}(\mathbf{z}|\mathcal{H}_m) = \int_{-\infty}^{\infty} f_{\mathbf{z}(u)|\Theta_c(u)}(\mathbf{z}|\phi, \mathcal{H}_m) f_{\Theta_c(u)}(\phi) d\phi
$$

$$
f_{\Theta_c(u)}(\phi) = f_{\Theta_c(u)}(\phi|\mathcal{H}_m)
$$

$$
= \frac{1}{2\pi} \phi \in [0, 2\pi)
$$

Let's evaluate this for the CT-Likelihood functional in AWGN

$$
L(\bm{r}|\mathcal{H}_m) = e^{-E_m/N_0} \exp\left(\frac{2}{N_0}\int_0^T \Re\left\{\bar{r}(t)s^*_m(t)\right\}dt\right)
$$

Phase Noncoherent Detection

Recall, the complex BB version of the likelihood

$$
L(\boldsymbol{r}|\mathcal{H}_m) = e^{-E_m/N_0} \exp\left(\frac{2}{N_0} \int_0^T \Re\left\{\bar{r}(t) s_m^*(t)\right\} dt\right)
$$

Modeling the unknown incoming phase offset

$$
s_m(t; \Theta_c(u)) = \Re \left\{ \bar{s}_m(t) \sqrt{2} e^{j(2\pi f_c t + \Theta_c(u))} \right\}
$$

$$
= \Re \left\{ \bar{s}_m(t) e^{j\Theta_c(u)} \sqrt{2} e^{j2\pi f_c t} \right\}
$$

$$
= \Re \left\{ \bar{s}_m(t; \Theta_c(u)) \sqrt{2} e^{j2\pi f_c t} \right\}
$$

$$
\bar{s}_m(t; \Theta_c(u)) = \bar{s}_m(t) e^{j\Theta_c(u)}
$$

Phase Noncoherent Detection $\operatorname{\mathsf{oncoher}}\nolimits$ *N*⁰ \overline{a} r **erent Detection** z
Z 2º

$$
L(r|\mathcal{H}_m) = e^{-E_m/N_0} \int_0^{2\pi} \exp\left(\frac{2}{N_0} \int_0^T \Re\left\{\bar{r}(t)s_m^*(t)e^{-j\phi}\right\} dt\right) \frac{d\phi}{2\pi}
$$

$$
= e^{-E_m/N_0} \frac{1}{2\pi} \int_0^{2\pi} \exp\left(\frac{2}{N_0}|\bar{r}_m|\cos(\phi - \angle \bar{r}_m)\right) d\phi
$$

$$
= e^{-E_m/N_0} \frac{1}{2\pi} \int_{2\pi} e^{\frac{2}{N_0}|\bar{r}_m|\cos(\psi)d\psi}
$$

define r_m

$$
I_0(x) = \frac{1}{2\pi} \int_{2\pi} e^{x \cos \phi} d\phi
$$

ار
2≛ا Average Likelihood Functional, phase noncoherent AWGN *^I*0(*x*) = ¹ Z

$$
L(\boldsymbol{r}|\mathcal{H}_m) = e^{-E_m/N_0} I_0\left(\frac{2}{N_0}|\bar{r}_m|\right)
$$

Phase Noncoherent Detection

If signals are equal energy

$$
L(\bm{r}|\mathcal{H}_m) = e^{-E_m/N_0} I_0\left(\frac{2}{N_0}|\bar{r}_m|\right) \equiv I_0\left(\frac{2}{N_0}|\bar{r}_m|\right)
$$

^L(*r|Hm*) = *^eEm/N*⁰ *^I*⁰ 1
12 | Envelope Detector (equal energy noncoherent) *|r*¯*m|*

$$
\left(\max_{m} L(r|\mathcal{H}_m) \iff \max_{m} |\bar{r}_m|\right)
$$

Envelope Detector Processing

Envelope Detector (narrowband processing)

Example: Non-coherent BFSK

Noncoherent Binary (equal E) Performance = *P*(*E|H*1) *P*(*E*) *Performance*

Orthogonal Binary Noncoherent

$$
P(\mathcal{E}) = \frac{1}{2} \exp\left(\frac{-E}{2N_0}\right)
$$

 $=$ \overline{a} $\Pr\left\{ |\bar{r}_0(u)| < |\bar{r}_1(u)| |\mathcal{H}_0 \right\}.$ 2*N*⁰ $P(\mathcal{E}|\mathcal{H}_0) = \Pr\{|\bar{r}_0(u)| < |\bar{r}_1(u)||\mathcal{H}_0\}$

> r fra = Pr *{*Rayleigh rv *>* Rice rv*}*

b = $= P(\mathcal{E}|\mathcal{H}_1)$

$$
P(\mathcal{E}) = Q(a, b) - \frac{1}{2} \exp\left(\frac{-(a^2 + b^2)}{2}\right) I_0(ab)
$$

$$
a = \sqrt{\frac{E}{2N_0} \left(1 - \sqrt{1 - |\rho_c|^2}\right)}
$$

$$
b = \sqrt{\frac{E}{2N_0} \left(1 + \sqrt{1 - |\rho_c|^2}\right)}
$$

$$
Q(a, b) = \int_b^\infty x \exp\left(\frac{-(a^2 + b^2)}{2}\right) I_0(abx) dx
$$

 $\rho_c =$ 1 *E* \int_0^T $\overline{0}$ $\bar{s}_0(t)\bar{s}_1^*(t)dt$ *^P*(*E*) = Q(*a, b*) ¹ $E \int_0^{\infty} e^{-\theta (x^2 + 1)^2} dx$

> Marcum Q-function \overline{D}

Noncoherent Binary (equal E) Performance

Best non-coherent performance is for orthogonal
Phase Noncoherent Detector (signal basis)

complex BB correlation to basis signals are sufficient statistics

 ${\lbrace \bar{\phi}_i(t) \rbrace}_{i=1}^{\bar{D}}$ = Orthonomal basis for complex BB model

$$
\int_{\mathcal{T}} \bar{r}(t) \bar{s}_m^*(t) dt = \int_{\mathcal{T}} \bar{r}(t) \sum_{i=1}^{\bar{D}} \bar{S}_m^*(i) \bar{\phi}_i^*(t) dt
$$

$$
= \sum_{i=1}^{\bar{D}} \bar{S}_m^*(i) \int_{\mathcal{T}} \bar{r}(t) \bar{\phi}_i^*(t) dt
$$

 \bigcap \overline{I} $\langle \bar{r}(t) \bar{\phi}_i^*(t) dt \rangle = \text{ sufficient statistics}$

Phase Noncoherent Detector (signal basis) *l*etector

Differential Encoding of PSK

differential phase encoder

Can detect several ways

Differential PSK

• Coherent detection with differential decoding

- First do coherent MSK detection, then put hard symbol decisions through inverse of differential encoder
- *• "Differentially Coherent" detection (DPSK)*
	- Do phase noncoherent detection over two symbol times
- **• Optimal MAP detection**
	- Optimal processing decides by processing entire sequence
		- Viterbi or Forward-Backward Algorithm

Differentially Coherent Detection of DE-PSK *^z*¯*k*1(*u*) = ^p*Esejk*¹ *^ej*✓*c*(*u*) + ¯*wk*1(*u*)

noncoherent based on two symbols

$$
\mathcal{H}_m: \begin{bmatrix} \bar{z}_k(u) \\ \bar{z}_{k-1}(u) \end{bmatrix} = \sqrt{E_s} \begin{bmatrix} e^{j\phi_k} \\ e^{j\phi_{k-1}} \end{bmatrix} e^{j\theta_c(u)} + \begin{bmatrix} \bar{w}_k(u) \\ \bar{w}_{k-1}(u) \end{bmatrix} \quad \theta_k = \phi_k - \phi_{k-1} = \frac{2\pi}{M}m
$$

$$
\mathcal{H}_m: \quad \bar{\mathbf{z}}(u) = \bar{\mathbf{y}}_m e^{j\theta_c(u)} + \bar{\mathbf{w}}(u)
$$

$$
\max_{m} \left| \bar{\mathbf{y}}_{m}^{\dagger} \bar{\mathbf{z}} \right|^{2} \iff \max_{m} \left| e^{-j\phi_{k-1}} \left[e^{-\theta_{k}} \quad 1 \right] \left[\begin{array}{c} \bar{z}_{k} \\ \bar{z}_{k-1} \end{array} \right] \right|^{2}
$$

Differentially-Coherent PSK Demod

$$
\iff \max_{m} \left| \bar{z}_k e^{-j\theta_k} + \bar{z}_{k-1} \right|^2
$$

$$
\iff \max_{\theta_k \in \left\{\frac{2\pi}{M}m\right\}} \Re\left\{\bar{z}_k \bar{z}_{k-1}^* e^{-\theta_k}\right\}
$$
\n
$$
\iff \min_{\theta_k \in \left\{\frac{2\pi}{M}m\right\}} \left| \angle(\bar{z}_k \bar{z}_{k-1}^*) - \theta_k \right|
$$

Differentially Coherent Detection of DE-PSK

$$
\mathcal{H}_m: \begin{bmatrix} \bar{z}_k(u) \\ \bar{z}_{k-1}(u) \end{bmatrix} = \sqrt{E_s} \begin{bmatrix} e^{j\phi_k} \\ e^{j\phi_{k-1}} \end{bmatrix} e^{j\theta_c(u)} + \begin{bmatrix} \bar{w}_k(u) \\ \bar{w}_{k-1}(u) \end{bmatrix} \quad \theta_k = \phi_k - \phi_{k-1} = \frac{2\pi}{M}m
$$

Performance of DC-BPSK

$$
\mathcal{H}_0: \begin{bmatrix} \bar{z}_k(u) \\ \bar{z}_{k-1}(u) \end{bmatrix} = \sqrt{E_s} \begin{bmatrix} +1 \\ +1 \end{bmatrix} e^{j\theta_c(u)} + \begin{bmatrix} \bar{w}_k(u) \\ \bar{w}_{k-1}(u) \end{bmatrix}
$$

$$
\mathcal{H}_1: \begin{bmatrix} \bar{z}_k(u) \\ \bar{w}_k(u) \end{bmatrix} = \sqrt{E_s} \begin{bmatrix} -1 \\ -1 \end{bmatrix} e^{j\theta_c(u)} + \begin{bmatrix} \bar{w}_k(u) \\ \bar{w}_k(u) \end{bmatrix}
$$

$$
\mathcal{H}_1: \begin{bmatrix} \bar{z}_k(u) \\ \bar{z}_{k-1}(u) \end{bmatrix} = \sqrt{E_s} \begin{bmatrix} -1 \\ +1 \end{bmatrix} e^{j\theta_c(u)} + \begin{bmatrix} \bar{w}_k(u) \\ \bar{w}_{k-1}(u) \end{bmatrix}
$$

Binary, orthogonal, noncoherent:

$$
P(\mathcal{E}) = \frac{1}{2} \exp\left(\frac{-(2E)}{2N_0}\right) = \frac{1}{2} \exp\left(\frac{-E_b}{N_0}\right)
$$

DBPSK Performance

6 \overline{e} <u>ا</u>
ا *z*¯*k*(*u*) erformance ca \overline{Y} \mathbf{y} ormance ca
10n-orthog \mathcal{E} h oounded with f
. $\overline{\mathsf{in}}$ $\overline{ }$ unded with r
ary noncohe \overline{a} DC-MPSK Performance can be bounded with PWerror given by non-orthogonal, binary noncoherent

Comparison of Binary Signaling/Detection Methods

DBPSK is a simple way to approach coherent BPSK without a phase reference

Soft-out Demapper (SOMAP, Soft-dempper)

- Using M-ary modulation with q bits labeling each symbol
- Have focused on MAP symbol detection
	- Selecting the MAP symbol implied a decision on the q bit labels
- We will now consider the MAP rule for deciding each bit
	- other bits are viewed as nuisance parameters and form the average likelihood

Motivation: Bit-Interleaved Coded Modulation (BICM)

BICM with Iterative Decoding/Demod

In general, soft-demapper should take in a priori soft decision information on dj as well as channel likelihoods

Soft-out Demapper (SOMAP, Soft-demapper)

$$
f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|d_j) = \sum_{\bar{\mathbf{d}}_j} f_{\mathbf{z}(u)|\mathbf{d}(u)}(\mathbf{z}|\bar{\mathbf{d}}_j, d_j) p_{\bar{\mathbf{d}}_j(u)}(\bar{\mathbf{d}}_j)
$$

$$
= \sum_{\bar{\mathbf{d}}_j} \left[f_{\mathbf{z}(u)|\mathbf{x}(u)}(\mathbf{z}|\mathbf{x}(\mathbf{d})) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]
$$

$$
\equiv \sum_{\mathbf{\bar{d}}_j} \left[\exp\left(\frac{-\|\mathbf{z} - \mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]
$$

if AWGN channel

$$
\equiv \sum_{\bar{\mathbf{d}}_j} \exp\left(\frac{-\|\mathbf{z} - \mathbf{x}(\mathbf{d})\|^2}{N_0}\right)
$$

if d's are a priori uniform

 $\bar{\mathbf{d}}_j = \{d_i\}_{i \neq j}$

nuisance parameters in this context

Soft-out Demapper (SOMAP, Soft-demapper)

for each bit location, we average over the the subset of signals with a 0 in location j, then over all points with 1 in location j

Soft-out Demapper (SOMAP, Soft-demapper) ADD^o d¯*j* $er~(SOMAPS)$ *N*⁰

$$
\frac{f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)}{f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)} = \frac{\sum_{\bar{\mathbf{d}}_j,d_j=1} \left[\exp\left(\frac{-\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]}{\sum_{\bar{\mathbf{d}}_j,d_j=0} \left[\exp\left(\frac{-\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]}
$$

 \overline{a} *y* erage likeliho \mathbb{Z} ati_' d¯*j* or bit dj — soft-decision *ⁱ*6=*^j ^pdi*(*u*)(*di*) average likelihood ratio for bit dj — soft-decision sent to decoder

$$
f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)p_{d_j(u)}(1) \underset{\mathcal{H}_0}{\geq} f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)p_{d_j(u)}(0)
$$

This is the MAP bit decision rule for bit dj

Decision Format vs. Optimality Criterion *f*z(*u*)*|d^j* (*u*)(z*|*1)*pd^j* (*u*)(1) *H*1 *> < f*z(*u*)*|d^j* (*u*)(z*|*0)*pd^j* (*u*)(0)

$$
\left\{ f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)p_{d_j(u)}(1) \sum_{\mathcal{H}_0}^{q_1} f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)p_{d_j(u)}(0) \right\}_{j=0,1,\ldots,q-1}
$$

The q MAP bit decision rules imply an M-ary decision rule *>*

This is the M-ary Bayes rule with $C(i,j)$ = number of bit label differences *f*z(*u*)*|d^j* (*u*)(z*|*1)*pd^j* (*u*)(1) *<* $(I,$ *f*z(*u*)*|d^j* (*u*)(z*|*0)*pd^j* (*u*)(0)

minimizes the average number of bit errors, or Pb *f*z(*u*)*|d^j* (*u*)(z*|*1)*pd^j* (*u*)(1) *> F* (*u*)
F (*u*) *d*

$$
\max_{m \in \{0,1,\ldots M-1\}} \left[f_{\mathbf{z}(u)|\mathbf{x}(u)}(\mathbf{z}|\mathbf{s}_m) \prod_{i=0}^{q-1} p_{d_i(u)}(d_i(m)) \right]
$$

The one M-ary MAP symbol decision rule implies q bit decision rules — **what are these?**

Decision Format vs. Optimality Criterion

$$
g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|d_j) = \max_{\bar{\mathbf{d}}_j} f_{\mathbf{z}(u)|\mathbf{d}(u)}(\mathbf{z}|\bar{\mathbf{d}}_j, d_j) p_{\bar{\mathbf{d}}_j(u)}(\bar{\mathbf{d}}_j)
$$

$$
= \max_{\bar{\mathbf{d}}_j} \left[f_{\mathbf{z}(u)|\mathbf{x}(u)}(\mathbf{z}|\mathbf{x}(\mathbf{d})) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]
$$

$$
\equiv \max_{\mathbf{\vec{d}}_j} \left[\exp\left(\frac{-\|\mathbf{z} - \mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]
$$

$$
\equiv \max_{\bar{\mathbf{d}}_j} \exp\left(\frac{-\|\mathbf{z} - \mathbf{x}(\mathbf{d})\|^2}{N_0}\right)
$$

$$
\left\{g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)p_{d_j(u)}(1) \sum_{i=0}^{n-1} g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)p_{d_j(u)}(0)\right\}_{j=0,1,\ldots q-1}
$$

MAP M-ary symbol decision rule expressed as q bitlevel decisions

Decision Format vs. Optimality Criterion MI[*di*] = ln

$$
\frac{g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)}{g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)} = \frac{\max_{\mathbf{\bar{d}}_j,d_j=1} \left[\exp\left(\frac{-\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]}{\max_{\mathbf{\bar{d}}_j,d_j=0} \left[\exp\left(\frac{-\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]}
$$

generalized likelihood ratio for bit dj — soft-decision sent to decoder

SOMAP processing in AWGN \overline{SOMAP} processing in A

$$
\text{sum-product SOMAP} \quad \frac{f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)}{f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)} = \frac{\sum_{\bar{\mathbf{d}}_j,d_j=1} \left[\exp\left(\frac{-\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]}{\sum_{\bar{\mathbf{d}}_j,d_j=0} \left[\exp\left(\frac{-\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]}
$$

$$
\text{max-product SOMAP} \quad \frac{g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)}{g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)} = \frac{\max_{\bar{\mathbf{d}}_j,d_j=1} \left[\exp\left(\frac{-\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]}{\max_{\bar{\mathbf{d}}_j,d_j=0} \left[\exp\left(\frac{-\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N_0}\right) \prod_{i \neq j} p_{d_i(u)}(d_i) \right]}
$$

both of these can be implemented in the metric domain (-ln(.))

Metric Domain Processing for Max-Product

$$
-\ln\left(\max_{m} p_m\right) = -\max_{m} \left[-\ln(p_m)\right] = \min_{m} \left[-\ln(p_m)\right]
$$

$$
-\ln\left(\frac{g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)}{g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)}\right) = \min_{\bar{\mathbf{d}}_j,d_j=1} \left(\text{MI}[\mathbf{x}(\mathbf{d})] + \sum_{i \neq j} \text{MI}[d_i] \right) - \min_{\bar{\mathbf{d}}_j,d_j=0} \left(\text{MI}[\mathbf{x}(\mathbf{d})] + \sum_{i \neq j} \text{MI}[d_i] \right)
$$

min-sum SOMAP (metric domain implementation of max-product) $\overline{}$ $\lim_{x \to 0}$ of max S^{uuc}

 $\text{MI}[\mathbf{x}(\mathbf{d})] = -\ln \big[f_{\mathbf{z}(u)|\mathbf{x}(u)}(\mathbf{z}|\mathbf{x}(\mathbf{d})) \big]$

$$
= \frac{\|\mathbf{z} - \mathbf{x}(\mathbf{d})\|^2}{N_0} \quad \text{(AWGN)}
$$

 $\text{MI}[d_i] = -\ln (p_{d_i(u)}(d_i))$

Metric Domain Processing for Sum-Product

$$
-\ln\left(\sum_{m}p_{m}\right) = \min_{m}^{*}\left[-\ln(p_{m})\right]
$$

metric domain averaging

$$
\min^*(m_1, m_2) = -\ln(e^{-m_1} + e^{-m_2})
$$

$$
\min_i^* m_i = -\ln\left(\sum_i e^{-m_i}\right)
$$

simple implantation as a pairwise operation:

$$
\min^*(m_1, m_2) = -\ln(e^{-m_1} + e^{-m_2})
$$

=
$$
\min(m_1, m_2) - \ln(1 + e^{-|m_1 - m_2|})
$$

$$
\min^*(m_1, m_2, m_3) = \min^*(\min^*(m_1, m_2), m_3)
$$

Metric Domain Processing for Max-Product

$$
-\ln\left(\frac{f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)}{f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)}\right)=\min_{\bar{\mathbf{d}}_j,d_j=1}^*\left(\mathrm{MI}[\mathbf{x}(\mathbf{d})]+\sum_{i\neq j}\mathrm{MI}[d_i]\right)-\min_{\bar{\mathbf{d}}_j,d_j=0}^*\left(\mathrm{MI}[\mathbf{x}(\mathbf{d})]+\sum_{i\neq j}\mathrm{MI}[d_i]\right)
$$

 $min*$ -sum SOMAP (metric domain implementation of sum-product) ^k^z ^x(d)k²

$$
\mathrm{MI}[\mathbf{x}(\mathbf{d})] = -\ln \big[f_{\mathbf{z}(u)|\mathbf{x}(u)}(\mathbf{z}|\mathbf{x}(\mathbf{d})) \big]
$$

$$
= \frac{\|\mathbf{z} - \mathbf{x}(\mathbf{d})\|^2}{N_0} \quad \text{(AWGN)}
$$

 $\text{MI}[d_i] = -\ln (p_{d_i(u)}(d_i))$

(replace min with min*)

 $\text{MI}[\mathbf{x}(\mathbf{d})] = -\ln \big[f_{\mathbf{z}(u)|\mathbf{x}(u)}(\mathbf{z}|\mathbf{x}(\mathbf{d})) \big]$

1. Compute the configure metrics by *combining* incoming metrics

$$
= \frac{d\mathbf{a} - \mathbf{a}(\mathbf{a})}{N_0} \quad \text{(AWGN)}
$$

$$
\text{MI}[d_i] = -\ln\left(p_{d_i(u)}(d_i)\right)
$$

 $=\frac{\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N}$

$$
M[Config = m] = MI[s_m] + \sum_{i=0}^{q-1} MI[d_i^{(m)}]
$$

MSM[*dⁱ* = 1] = min *^m*:*di*=1 M[Config = *^m*] 2. *Marginalize* the configuration metric to get the marginal soft decision information *i*=0 δ ⁻¹

"Intrinsic" (soft) information

$$
MSM[d_i = 1] = \min_{m:d_i=1} M[Config = m]
$$

MSM[d_i = 0] = min M[Config = m]

threshold these for best local decisions — i.e., MAP symbol/ bit-sequence

 α to "extrinsic format" — i.e.. likelihoods t to extrinsic format — i.e., likelinoods 3. Convert to "extrinsic format" — i.e., likelihoods μ - Direct to extrinsic format — i.e., like
d

 $\text{MSM}[d_i = 0] = \min_{m:d_i=0} \text{M}[\text{Config} = m]$

"Extrinsic" (soft) information

$$
MO[d_i = 1] = MSM[d_i = 1] - MI[d_i = 1]
$$

$$
MO[d_i = 0] = MSM[d_i = 0] - MI[d_i = 0]
$$

MSM[*di*] = MSM[*dⁱ* = 1] MSM[*dⁱ* = 0]

pass these to the decoder as soft decisions

 $\text{MI}[\mathbf{x}(\mathbf{d})] = -\ln \big[f_{\mathbf{z}(u)|\mathbf{x}(u)}(\mathbf{z}|\mathbf{x}(\mathbf{d})) \big]$

 $=\frac{\|{\bf z} - {\bf x}({\bf d})\|^2}{N}$ *N*⁰

(AWGN)

 $\text{MI}[d_i] = -\ln (p_{d_i(u)}(d_i))$ Modulation De-mapper . . . \longleftarrow MI[x] $\mathrm{MI}[d_0] \longrightarrow$ $\mathrm{MI}[d_1]$ – $\text{MI}[d_{q-1}] \longrightarrow$ $MO[d_{q-1}] \leftarrow$ $MO[d_1]$ \leftarrow $MO[d_0]$ \longrightarrow

 $\text{MI}[\mathbf{x}(\mathbf{d})] = -\ln \big[f_{\mathbf{z}(u)|\mathbf{x}(u)}(\mathbf{z}|\mathbf{x}(\mathbf{d})) \big]$

 $=\frac{\|\mathbf{z}-\mathbf{x}(\mathbf{d})\|^2}{N}$

1. Compute the configure metrics by *combining* incoming metrics

$$
= \frac{d\mathbf{a} - \mathbf{a}(\mathbf{a})}{N_0} \quad \text{(AWGN)}
$$

$$
\text{MI}[d_i] = -\ln\left(p_{d_i(u)}(d_i)\right)
$$

$$
\mathbf{M}[\text{Config} = m] = \mathbf{M}\mathbf{I}[\mathbf{s}_m] + \sum_{i=0}^{q-1} \mathbf{M}\mathbf{I}[d_i^{(m)}]
$$

MSM[*dⁱ* = 1] = min *^m*:*di*=1 M[Config = *^m*] 2. *Marginalize* the configuration metric to get the marginal soft decision information *i*=0 **In the configuration** meand to go *q*1 (*m*)

 ${}^*{\rm M}[\text{Config} = m]$

⇤M[Config = *m*]

"Intrinsic" (soft) information

$$
MS^*M[d_i = 1] = \min_{m:d_i=1}^*M[Config = m]
$$

threshold these for best local decisions — i.e., MAP bit

 $m:d_i=0$

m:*di*=1

 $\text{MS}^*\text{M}[d_i=0]=\ \ \min_i$

MS⇤M[*dⁱ* = 1] = min

3. Convert to "extrinsic format" — i.e., likelihoods *m*:*di*=0 format" — i.e., lik

"Extrinsic" (soft) information

$$
MO[d_i = 1] = MS^*M[d_i = 1] - MI[d_i = 1]
$$

$$
MO[d_i = 0] = MS^*M[d_i = 0] - MI[d_i = 0]
$$

pass these to the decoder as soft decisions

Can work with the Negative Log-Likelihood Ratios (NLLRs) instead \int an work with the Negat

$$
\overline{\mathrm{MI}}[d_i] = \mathrm{MI}[d_i] - \mathrm{MI}[d_i = 0]
$$

$$
\overline{\text{MO}}[d_i] = \text{MO}[d_i] - \text{MO}[d_i = 0]
$$

⁼ ln *^f*z(*u*)*|d^j* (*u*)(z*|*1) *^f*z(*u*)*|d^j* (*u*)(z*|*0)! Can subtract any constant from metrics

SOMAP Processing — Normalized Metrics

Can always represent metrics/probabilities on M-ary variables by M-1 numbers through normalization

$$
\overline{\mathrm{MI}}[d_i] = \mathrm{MI}[d_i] - \mathrm{MI}[d_i = 0]
$$

$$
\overline{\mathrm{MI}}[d_i = 1] = \mathrm{MI}[d_i = 1] - \mathrm{MI}[d_i = 0]
$$

$$
= -\ln\left[\frac{p(d_i=1)}{p(d_i=0)}\right]
$$

 $\mathrm{MI}[d_i=0]=0$ "zeros are free"

(see spreadsheet example)

$$
\overline{\text{MO}}[d_i] = \text{MO}[d_i] - \text{MO}[d_i = 0]
$$

$$
\overline{\text{MO}}[d_i = 1] = \text{MO}[d_i = 1] - \text{MO}[d_i = 0]
$$

$$
= -\ln\left(\frac{f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)}{f_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)}\right) \quad (\text{min}^*-\text{sum})
$$

$$
= -\ln\left(\frac{g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|1)}{g_{\mathbf{z}(u)|d_j(u)}(\mathbf{z}|0)}\right) \quad (\text{min-sum})
$$

 $\overline{\mathrm{MI}}[d_i = 0] = 0$ "zeros are free"

 $\text{MI}[d_i] = \text{MI}[d_i] - \text{MI}[d_i = 0]$ $\overline{\text{MI}}[d\cdot] = \text{MI}[d\cdot] = \text{MI}[d\cdot] = 0$

 $\overline{\mathrm{MI}}[d_i = 1] = \mathrm{MI}[d_i = 1] - \mathrm{MI}[d_i = 0]$

I often abuse this notation and use the first to represent the second — i.e., since once of the two normalized metrics is zero by definition

$$
\overline{\mathrm{MO}}[d_i] = \mathrm{MO}[d_i] - \mathrm{MO}[d_i = 0]
$$

$$
\overline{\text{MO}}[d_i = 1] = \text{MO}[d_i = 1] - \text{MO}[d_i = 0]
$$

for equal a priori probability on the bits — e.g., first activation and/ or non-iterative BICM

$$
\begin{aligned}\n\overline{\text{min-sum:}} & \overline{\text{MO}}[d_j] = \min_{m:d_j=1} \frac{\|\mathbf{z} - \mathbf{s}_m\|^2}{N_0} - \min_{m:d_j=0} \frac{\|\mathbf{z} - \mathbf{s}_m\|^2}{N_0} \\
\overline{\text{min*-sum:}} & \overline{\text{MO}}[d_j] = \min_{m:d_j=1} \frac{\|\mathbf{z} - \mathbf{s}_m\|^2}{N_0} - \min_{m:d_j=0} \frac{\|\mathbf{z} - \mathbf{s}_m\|^2}{N_0} \\
\end{aligned}
$$

for equal a priori probability on the bits — e.g., first activation and/ or non-iterative BICM

SOMAP Processing is Special Case of SISO

- General Soft-in/Soft-out (SISO) processing
	- Digital variables (e.g., inputs/outputs) associated with a local system/ constraint/code
		- Finite number of configurations
	- Combine incoming marginal soft information (e.g., sum MI's) to compute a configuration metric for each configuration
	- Marginalize over configuration metrics consistent with each value of each digital variable to produce updated marginal soft information (MO's)
- This forms the basis of all modern coding i.e., it is the basis of iterative decoding
	- Modern coding: decode local codes in SISO manner, exchange soft information, and iterate